Stat 515: Introduction to Statistics

Chapter 6

Recall Definitions from Ch 2

- **Statistic**: numerical summary of a sample
 - Mean(\bar{x}), proportion(\hat{p}), median, mode, standard deviation(s), variance(s^2), Q1, Q3, IQR, etc.
 - We use US alphabet letters to denote these
- **Parameter**: numerical summary of a population
 - Mean(μ_x), proportion(ρ), median, mode, standard deviation(σ), variance(σ^2), Q1, Q3, IQR, etc.
 - We usually don't know these values
 - We use Greek letters to denote these

- A sampling distribution is the probability distribution that specifies probabilities for the possible values of the sample mean or proportion.
- A sampling distribution is a special case of a probability distribution where the outcome of an experiment that we are interested in is a sample statistic such as a sample proportion(p̂) or sample mean (x̄)
 - It's the same as what we were doing before, but now instead of singular observations we're looking at groups

- This is confusing.
 - Remember, before we talked about events and random variables in n trials
 - Now, we're talking about m groups of n trials which yield m sample means or m sample proportions

•
$$\overline{x_i} = \frac{\sum x}{n}$$
 for $i = 1, 2, ..., m$
• $\widehat{p_i} = \frac{x}{n}$ for $i = 1, 2, ..., m$

• Variable: Gender of Students

- Before, we measured individuals:

Bob Joe Jane Kate Bill Amy Ray Chris Jess Male Male Female Female Male Female Male Male Female - Now, we have one measurement across groups:



• Variable: Heights of Americans

- Before, we measured individuals:



– Now, we have one measurement across groups:



Sampling Distribution - Graphs

 Sample vs. Population: the sampling distribution is narrower than the population because grouping the data reduces the variation; pay attention to the standard error equations



- This first sampling distribution we'll talk about is the sampling distribution for the sample proportion \hat{p} .
- The idea is that there is some true population proportion out there, ρ, but in most cases it isn't feasible to know it
 - We may not have enough time or money to poll the population
 - It may be infeasible to get a population measure

• We look at sample proportions, \hat{p} , the proportion of observations in our sample that have a certain characteristic among our sample

– Think "x out of n" then
$$\widehat{p}=rac{x}{n}$$

 We've looked at this before in the descriptive statistics but now we're going to talk about all possible sample proportions from repeated random samples from the population and their distribution (mean and standard deviation)

- Before we had categorical observations: $x_1, x_2, x_3, \dots, x_n$
 - We would summarize all x's with one sample proportion, one \widehat{p}

•
$$\hat{p} = \frac{\text{number of x with desired trait}}{\text{total sample size}}$$

= the proportion of our sample with the desired trait

- Now we have m groups of n subjects with categorical observations: $\{x_{1,1}, x_{1,2}, x_{1,3}, \dots, x_{1,n}\}, \{x_{2,1}, x_{2,2}, x_{2,3}, \dots, x_{2,n}\}, \dots, \{x_{m,1}, x_{m,2}, x_{m,3}, \dots, x_{m,n}\}$
- Now, we find summary statistics for each group $\widehat{p_1}, \widehat{p_2}, \widehat{p_3}, \widehat{p_4}, \dots, \widehat{p_m}$
 - We have m sample proportions , one \hat{p} for each group

$$- \widehat{p_1} = \frac{\text{number of x with desired trait in group 1}}{\text{total sample size of group 1}}$$
$$- \widehat{p_2} = \frac{\text{number of x with desired trait in group 2}}{\text{total sample size of group 2}} \dots$$
$$- \widehat{p_m} = \frac{\text{number of x with desired trait in group } m}{\text{total sample size of group m}}$$

- You could think of each group as a barrel and we're only interested in the proportion of each barrel; we are no longer interested in the individual responses like we might have been before
- The example below shows how we could summarize 40 observations by splitting them into four representative sample proportions

 $\widehat{p_1} \qquad \widehat{p_2} \qquad \widehat{p_3} \qquad \widehat{p_4}$ $x_1, x_2, \dots, x_{10} \qquad x_{11}, x_{12}, \dots, x_{20} \qquad x_{21}, x_{22}, \dots, x_{30} \qquad x_{31}, x_{32}, \dots, x_{40}$

- The mean of the sampling distribution for a sample proportion will always equal the population proportion: $\mu_{\widehat{p}} = \rho$
 - Even though we know the mean is the population proportion, we note that some \hat{p} will be lower and some will be higher

- Think about it this way:
 - Q: If the population proportion of females in the United States is 51% what would you expect the number of females to be in a random sample of 100 Americans?
 - A: 51%, or 51 of 100, is our best guess; think of the binomial expectation.
- Later, we'll do this the other way around and we will call p̂ the point estimate for ρ since it's our best guess for the population proportion if we don't know it

• The **standard error**, the standard deviation of all possible sample proportions, is:

$$\sigma_{\widehat{p}} = \sqrt{\frac{\rho(1-\rho)}{n}}$$
$$= St. Dev(\widehat{p_1}, \widehat{p_2}, \widehat{p_3}, \widehat{p_4}, \dots, \widehat{p_m})$$

- Think about it this way:
 - **Q:** If our best guess for ρ is \hat{p} we need a **measure of reliability** for our estimate
 - A: We'll talk more about this later, but our standard error calculator is a big part of this

• Recall:
$$\sigma_{\widehat{p}} = \sqrt{rac{
ho(1-
ho)}{n}}$$

- Later, in the case we don't know ${\bf \rho}$ we're estimating it with our point estimate \hat{p}

– Consider:

$$\int \frac{\hat{p}(1-\hat{p})}{n}$$

- $\mu_{\widehat{p}} = \rho$
 - Even though we know the mean is the population proportion, we note that some \hat{p} will be lower and some will be higher

•
$$\sigma_{\widehat{p}} = \sqrt{\frac{\rho(1-\rho)}{n}}$$

- Aside:
 - What if we increase n?
 - The standard deviation shrinks
 - What if we decrease n?
 - The standard deviation grows

 Now that we know the mean and standard deviation of the sample proportions we can calculate z-scores to find some probabilities associated with sample proportions just like we did before.

$$\mu_{\hat{p}} = \rho$$

$$\sigma_{\hat{p}} = \sqrt{\frac{\rho(1-\rho)}{n}}$$

$$z = \frac{observation-mean}{st.dev} = \frac{\hat{p}-\mu_{\hat{p}}}{\sigma_{\hat{p}}} = \frac{\hat{p}-\rho}{\sqrt{\frac{\rho(1-\rho)}{n}}}$$

$$P(\hat{p} > c) = 1 - P\left(z < \frac{c - \mu_{\hat{p}}}{\sigma_{\hat{p}}}\right) = 1 - P\left(z < \frac{c - \rho}{\sqrt{\frac{\rho(1 - \rho)}{n}}}\right)$$
$$P(\hat{p} < c) = P\left(z < \frac{c - \mu_{\hat{p}}}{\sigma_{\hat{p}}}\right) = P\left(z < \frac{c - \rho}{\sqrt{\frac{\rho(1 - \rho)}{n}}}\right)$$

$$P(c_1 < \hat{p} < c_2) = P\left(z < \frac{c_2 - \mu_{\hat{p}}}{\sigma_{\hat{p}}}\right) - P\left(z < \frac{c_1 - \mu_{\hat{p}}}{\sigma_{\hat{p}}}\right)$$
$$= P\left(z < \frac{c_2 - \rho}{\sqrt{\frac{\rho(1 - \rho)}{n}}}\right) - P\left(z < \frac{c_1 - \rho}{\sqrt{\frac{\rho(1 - \rho)}{n}}}\right)$$

- Say, we know that 16% of Americans approve of Congress (Gallup).
- What is the sampling distribution of the sample proportion of Americans that approve of Congress for n=100?
 - Note, we aren't interested in the yes or no's individually but the proportion among the ten
 - Here, X=the proportion of the one hundred
 Americans in each group

- Say, we know that 16% of Americans approve of Congress (Gallup).
- What is the sampling distribution of the sample proportion of Americans that approve of Congress for n=100?

- n = sample size = **sample size of one hundred**= 100
- p = population proportion = **16%** =.16

• Let's find the sampling distribution mean:

• The mean of all sample proportions of n=100 = $\mu_{\hat{p}} = \rho = 16\% = .16$

- Some \hat{p} will be lower and some will be higher but the mean of all sample proportions of n=100 will be .6

- Let's find the sampling distribution st. error:
- The st. deviation of all sample proportions of n=100

Standard Error =
$$\sigma_{\hat{p}} = \sqrt{\frac{\rho(1-\rho)}{n}}$$

= $\sqrt{\frac{.16(1-.16)}{100}} = .0367$

• Let's find the sampling distribution :

$$\mu_{\hat{p}} = \rho = 16\% = .16$$

$$\sigma_{\hat{p}} = \sqrt{\frac{\rho(1-\rho)}{n}} = \sqrt{\frac{.16(1-.16)}{100}} = .0367$$

• The probability that **most**, of our sample of n=100, approve of Congress:

$$P(\hat{p} > .5) = 1 - P(\hat{p} < .5)$$

= 1 - pnorm(.5, .16, .0367)
 ≈ 0

Z-table:

$$P(\hat{p} > .5) = P\left(z > \frac{.5 - .16}{.0367}\right) = P(z > 9.26)$$
$$= 1 - P(z \le 9.26) \approx 1 - 1$$
$$= 0$$

Note: the z-table is a less accurate approximation than R

• The probability that **less than 10%**, of our sample of n=100, approve of Congress:

<u>R</u>

$$P(\hat{p} < .1) = pnorm(.1, .16, .0367) \\= .0510$$

Z-table:

$$P(\hat{p} < .1) = P\left(z < \frac{.1 - .16}{.0367}\right) = P(Z < -1.63)$$
$$= .0516$$

Note: the z-table is a less accurate approximation than R

- The probability that between 5 and 19 percent, of our sample of n=100, approve of Congress:
- $\frac{R}{P(.05 < \hat{p} < .19)} = P(\hat{p} < .19) P(\hat{p} < .05)$ = pnorm(.19, .16, .0367) - pnorm(.05, .16, .0367) = .7917991

<u>Z-table:</u>

$$P(.05 < \hat{p} < .19) = P(\hat{p} < .19) - P(\hat{p} < .05)$$

= $P\left(z < \frac{.19 - .16}{.0367}\right) - P\left(z < \frac{.05 - .16}{.0367}\right)$
= $P(Z < .82) - P(Z < -3.00)$
= .7939 - .0013
= .7926

• <u>Note</u>: we had to assume normality of \hat{p} to use the Z-score transformation and R code I provided to solve the previous probabilities

 Later, we will see how we are able to make that assumption – unlocking all of the nice methodologies of the Normal distribution

- This second sampling distribution we'll talk about is the sampling distribution for the sample mean.
- The idea is that there is some true population mean out there, μ, but it might not be feasible to know it
 - We may not have enough time or money to poll the population
 - It may be infeasible to get a population measure

- Instead, we look at sample mean, \overline{x} , the mean of quantitative observations
- We've looked at this before in the descriptive statistics but now we're going to talk about all possible sample means from repeated random samples from our population

- Before we had quantitative observations: $x_1, x_2, x_3, \dots, x_n$
 - We would summarize all x's with one sample mean, one \overline{x}

•
$$\bar{x} = \frac{\text{the sum of } x's}{\text{the total sample size}} = \frac{\sum x}{n}$$

= the mean of the observations in our sample

- Now we have m groups of n subjects with categorical observations:
 {x_{1,1}, x_{1,2}, x_{1,3}, ..., x_{1,n}}, {x_{2,1}, x_{2,2}, x_{2,3}, ..., x_{2,n}}, ..., {x_{m,1}, x_{m,2}, x_{m,3}, ..., x_{m,n}}
- Now, we find summary statistics for each group

 $\overline{x_1}$, $\overline{x_2}$, $\overline{x_3}$, $\overline{x_4}$, ..., $\overline{x_m}$

– We have m sample means, one \bar{x} for each group

- You could think of each group as a barrel and we're only interested in the mean of each barrel; we are no longer interested in the individual responses
- The example below shows how we could summarize 40 observations, into four representative sample means



- The mean of the sampling distribution for a sample mean will always equal the population mean: $\mu_{\overline{x}} = \mu_x$
 - This is the mean of all possible sample means, but we note that some \overline{x} will be lower and some will be higher

• Think about it this way:

- Q: If the population mean of time Americans spend on social media is 100 minutes with a standard deviation of 25 minutes what would you expect the average time a sample of 35 Americans spent on social media?
- A: 100 minutes is our best guess.
- Later, we'll do this the other way around and we will call x̄ the point estimate for μ_x since it's our best guess for the population mean if we don't know it

• The standard error, the standard deviation of all possible sample means, is:

$$\sigma_{\overline{x}} = \frac{\sigma_x}{\sqrt{n}}$$

= St. Dev($\overline{x_1}, \overline{x_2}, \overline{x_3}, \overline{x_4}, \dots, \overline{x_m}$)
Sampling Distribution – Mean and SD

• Think about it this way:

- Q: If our best guess for μ is \bar{x} we need a **measure of** reliability for our estimate
- A: We'll talk more about this later, but our standard error calculator is a big part of this
- Later, in the case we don't know μ_x or σ_x we're estimating it with our **point estimate** \overline{x}

- Recall:
$$\sigma_{\overline{x}} = \frac{\sigma_x}{\sqrt{n}}$$

- Consider: $\frac{s_x}{\sqrt{n}}$ [Note: we estimate $\sigma_x = s_x$]

Sampling Distribution – Mean and SD

- $\mu_{\overline{x}} = mean \ of \ all \ sample \ means = \mu_x$
 - Even though we know the mean is the population mean, we note that some \bar{x} will be lower and some will be higher

•
$$\sigma_{\overline{x}} = \frac{\sigma_x}{\sqrt{n}}$$

- Aside:
 - What if we increase n?
 - The standard deviation shrinks
 - What if we decrease n?
 - The standard deviation grows

Sampling Distribution:

 Now that we know the mean and standard error of the sample means we can calculate zscores to find some probabilities associated with sample means just like we did before.

$$\mu_{\overline{x}} = \mu_{x}$$

$$\sigma_{\overline{x}} = \frac{\sigma_{x}}{\sqrt{n}}$$

$$z = \frac{observation - mean}{st. dev} = \frac{\overline{x} - \mu_{\overline{x}}}{\sigma_{\overline{x}}} = \frac{\overline{x} - \mu_{x}}{\frac{\sigma_{x}}{\sqrt{n}}}$$

Sampling Distribution:

$$P(\bar{x} > c) = 1 - P\left(z < \frac{c - \mu_{\bar{x}}}{\sigma_{\bar{x}}}\right) = 1 - P\left(z < \frac{c - \mu_{x}}{\frac{\sigma_{x}}{\sqrt{n}}}\right)$$
$$P(\bar{x} < c) = P\left(z < \frac{c - \mu_{\bar{x}}}{\sigma_{\bar{x}}}\right) = P\left(z < \frac{c - \mu_{x}}{\frac{\sigma_{x}}{\sqrt{n}}}\right)$$

$$P(c_1 < \bar{x} < c_2) = P\left(z < \frac{c_2 - \mu_{\bar{x}}}{\sigma_{\bar{x}}}\right) - P\left(z < \frac{c_1 - \mu_{\bar{x}}}{\sigma_{\bar{x}}}\right)$$
$$= P\left(z < \frac{c_2 - \mu_x}{\frac{\sigma_x}{\sqrt{n}}}\right) - P\left(z < \frac{c_1 - \mu_x}{\frac{\sigma_x}{\sqrt{n}}}\right)$$

- Say, we know that the average American spends 100 minutes on social media per day with a standard deviation of 25 minutes.
- What is the sampling distribution of the sample mean of time Americans spend on social media for n=35?
 - Note, we aren't interested in the individuals but the group of thirty five
 - Here, X=the proportion of the ten Americans in each group

- Say, we know that the average American spends 100 minutes on social media per day with a standard deviation of 25 minutes.
- What is the sampling distribution of the sample mean of time Americans spend on social media for n=35?
 - n = sample size = **sample size of thirty five** = 35
 - μ_{χ} = population mean = 100
 - σ_x = population standard deviation = 25

• Let's find the sampling distribution mean:

- The mean of all sample ,eams of n=35 = $\mu_{\bar{x}} = \mu_x = 100$
 - Some \overline{x} will be lower and some will be higher but the mean of all sample means of n=35 will be 100

- Let's find the sampling distribution st. error:
- The st. deviation of all sample means of n=35
 = Standard Error

$$= \sigma_{\overline{x}} = \frac{\sigma_x}{\sqrt{n}} = \frac{25}{\sqrt{35}} = 4.2258$$

• Let's find the sampling distribution :

$$\mu_{\bar{x}} = \mu_x = 100$$

$$\sigma_{\bar{x}} = \frac{\sigma_x}{\sqrt{n}} = \frac{25}{\sqrt{35}} = 4.2258$$

 The probability that a sample of n=35 spend more than two hours on social media on average:

$$\frac{R}{P(\bar{x} > 120)} = 1 - P(\bar{x} < 120)$$

= 1 - pnorm(120,100,4.2258)
= .000001107046

<u>Z-table:</u>

$$P(\bar{x} > 120) = P\left(z > \frac{120 - 100}{4.2258}\right) = P(z > 4.73)$$
$$= 1 - P(z < 4.73) \approx 1 - 1$$
$$= 0$$

Note: the z-table is a less accurate approximation than R

 The probability that a sample of n=35 spend less than one hour on social media on average:

<u>R</u>

$$P(\bar{x} < 60) = P(\bar{x} < 60)$$

= pnorm(60,100,4.2258)
= 1.458519 * 10⁻²¹

<u>Z-table:</u>

$$P(\bar{x} > 60) = P\left(z < \frac{60 - 100}{4.2258}\right) = P(Z < -9.47)$$
$$= P(Z < -9.47)$$
$$\approx 0$$

Note: the z-table is a less accurate approximation than R

- The probability that a sample of n=35 spend between 1 and 1.5 hours on social media on average:
- $\frac{R}{P(60 < \bar{x} < 90)} = P(\bar{x} < 90) P(\bar{x} < 60)$ = pnorm(90,100,4.2258) - pnorm(60,100,4.2258) = .008980629

<u>Z-table:</u>

$$P(60 < \bar{x} < 90) = P\left(\frac{90 - 100}{4.2258} < z < \frac{60 - 100}{4.2258}\right)$$
$$= P(Z < -2.37) - P(Z < -9.47)$$
$$\approx .0089 - 0$$
$$= 0$$

Note: the z-table is a less accurate approximation than R

• <u>Note</u>: we had to assume normality of \bar{x} to use the Z-score transformation and R code I provided to solve the previous probabilities

 Later, we will see how we are able to make that assumption – unlocking all of the nice methodologies of the Normal distribution

• (LLN 1) – As the sample size increases the sample estimates (\bar{x} or \hat{p}) approach the population values (μ or σ)

 (LLN 2) – As the number of trials increase the proportion of occurrences of any given outcome approaches the probability in the long run.

- 10 flips: 6 heads were flipped
 - Total proportion $=\frac{x}{n}=\frac{6}{10}=.60=60\%$ heads
- 10 more flips: 5 heads were flipped
 - Total proportion $=\frac{x}{n} = \frac{5+6}{10+10} = \frac{11}{20} = .55 = 55\%$ heads
- 10 more flips: 5 heads were flipped
 - Total proportion $=\frac{x}{n} = \frac{11+5}{20+10} = \frac{16}{30} = .5333 = 53.33\%$ heads
- 10 more flips: 3 heads were flipped
 - Total proportion $=\frac{x}{n} = \frac{16+3}{30+10} = \frac{19}{40} = .475 = 47.5\%$ heads
- 10 more flips: 6 heads were flipped

• Total proportion
$$=\frac{x}{n} = \frac{19+6}{40+10} = \frac{25}{50} = .5 = 50\%$$
 heads

• (**LLN**) – As the number of flips 80 increase the proportion of 0.0 oroportion heads approaches the probability of 0.4 seeing a heads, 0.2 P(heads)=.5,which is the red 0.0 line. 10 20 30 40

flips

50

- At first the proportion is all over the place – you can see the large spikes in the graph
- Importantly, we see that the proportion of coins that landed on heads levels off and gets closer and closer to 50%, the probability, which is where we expect it to go 'in the long run!'



Central Limit Theorem: Proportions

For random sampling with a large sample size
 n, the sampling distribution of the sample
 proportion is approximately a normal
 distribution

$$-n * p \ge 15 \text{ and } n * (1-p) \ge 15$$

• Introduction:

- <u>https://www.youtube.com/watch?v=Pujol1yC1_A</u>

Central Limit Theorem: Means

For random sampling with a large sample size
 n, the sampling distribution of the sample
 mean is approximately a normal distribution
 – For us, 30 is close enough to infinity

• Introduction:

– <u>https://www.youtube.com/watch?v=Pujol1yC1_A</u>

Central Limit Theorem: Means

- 1) For any population the sampling distribution of \bar{x} is bell shaped when the sample size n is large, when n is thirty or more
- 2) The sampling distribution of \bar{x} is bell-shaped when the population distribution is distribution is bell-shaped, regardless of sample size
- 3) We do not know the shape of the sampling distribution of \bar{x} if the sample size is small and the population distribution isn't bell-shaped

Central Limit Theorem

For any population the sampling distribution of \bar{x} is bell shaped when the sample size n is large, when n is thirty or more **Note:** for small sample size we can't say this.

Population

 \overline{x} when n=2 \overline{x} when n=30



Central Limit Theorem

The sampling distribution of x_{bar} is bell-shaped when the population distribution is distribution is bell-shaped, regardless of sample size



Chapter 6 Supplement (not in text book)

Central Limit Theorem

- Recall: The sampling distribution of x_{bar} is bell-shaped even for n=1 if X follows the normal distribution
- If a simple random sample is drawn from the population then $z = \frac{\bar{x} \mu_x}{\sigma_{\bar{x}}}$ follows the standard normal distribution
- The difficulty in this is that we rarely know the value of the parameters: μ_x or σ_x in $\sigma_{\bar{x}} = \frac{\sigma_x}{\sqrt{n}}$

Central Limit Theorem

- To combat this difficulty we will cover a few more named distributions and sampling distributions
 - 1. s^2 and the χ^2 (chi-squared) distribution
 - 2. \bar{x} , s^2 and the Student's t-distribution
 - 3. Two variances and the F distribution

• The chi-squared density can be defined as follows:

$$f_{x(x)} = \frac{1}{2^{\frac{n}{2}}\Gamma\left(\frac{n}{2}\right)} x^{\frac{n}{2}-1} e^{-\frac{x}{2}} I(x > 0)$$

- Mean = n
- Variance = 2n

χ^2 (chi-squared) distribution in R

• P(X = x) = 0 as the probability of any one value is always zero

- $P(X \le x) = \text{pchisq}(x, n)$
- $P(X \ge x) = 1 \text{pchisq}(x, n)$
- $P(x_1 < X < x_2) = pchisq(x_2, n) pchisq(x_1, n)$

n=1 4 CO 2 > <u>۲</u> **o** -2 0 4 6 8 10

n=1 4 CO 2 > × **o** -2 0 4 6 8 10



Х



• The chi-squared distribution can be defined as follows:

$$X^2 = z_1^2 + z_2^2 + \dots + z_n^2$$

• X^2 follows the chi-squared distribution with n degrees of freedom where z_i are independent random variables that follow the standard normal distribution

- Consider x₁, x₂, ... x_n independent from a normal distribution
- If we look at the formula for the sample variance from chapter two we see that we are summing $(x_i \bar{x})^2$ (n normal random variables)

• We're close to showing s^2 follows the χ^2 distribution but $(x_i - \bar{x})^2$ are normal – not the standard normal.

- Now, if we can write it as a sum of squared standard normal random variables we can say that s^2 follows the χ^2 distribution
 - This requires a "cute" trick



$$X_{n-1}^{2} = \left(\frac{(n-1)s^{2}}{\sigma_{x}^{2}}\right) = \sum \left(\frac{x_{i} - \bar{x}}{\sigma_{x}}\right)^{2}$$

- Again, consider $\left(\frac{x_i \bar{x}}{\sigma_x}\right)$ on the right hand side
- If we change \bar{x} to μ we have the usual $z = \frac{x_i \mu_x}{\sigma_x}$
- Note: We lose one degree of freedom, going from n to n-1, because we are using \bar{x} instead of μ
Still dealing with the difficulty of needing to know the population standard deviation for the Central limit theorem we talk about the tdistribution

• The t density can be defined as follows:

$$f_{x(x)} = \frac{\Gamma\left(\frac{n+1}{2}\right)}{\sqrt{\pi n}\Gamma\left(\frac{n}{2}\right)} \left(1 + \frac{x^2}{n}\right)^{-\frac{n+1}{2}} I(X \in \mathbb{R})$$

- Mean = 0
- Variance = $\frac{n}{n-2}$

Properties of the t-distribution

- 1. The t-distribution is different for different degrees of freedom
- 2. The t-distribution is centered and symmetric at 0
- 3. The area under the curve is 1 and ½ on either side of 0
- 4. The probability approaches 0 as we move away from 0
- 5. The t-distribution has fatter tails than the standard normal
- 6. As the sample size increases t gets close to z

• P(X = x) = 0 as the probability of any one value is always zero

•
$$P(X \le x) = pt(x, n)$$

- $P(X \ge x) = 1 pt(x, n)$
- $P(x_1 < X < x_2) = pt(x_2, n) pt(x_1, n)$

The t-distribution



• The t distribution can be defined as follows:

$$t_n = \frac{Z}{\sqrt{\frac{\chi_n^2}{n}}}$$

• t_n follows the t distribution with n degrees of freedom where Z follows the standard normal distribution and χ^2 is chi-squared and divided by its degrees of freedom

• Relating it back to s_x^2 and χ_n^2



• Relating it back to s_x^2 and χ_n^2

$$t_{n-1} = \frac{\left(\frac{\bar{x} - \mu_x}{\frac{\sigma_x}{\sqrt{n}}}\right)}{\sqrt{\frac{s_x^2}{\sigma_x^2}}} = \frac{\left(\frac{\bar{x} - \mu_x}{\frac{\sigma_x}{\sqrt{n}}}\right)}{\frac{s_x}{\sigma_x}} = \frac{\left(\frac{\bar{x} - \mu_x}{\frac{\sigma_x}{\sqrt{n}}}\right)}{\frac{s_x}{\sigma_x}} \frac{\sigma_x}{\frac{\sigma_x}{\sigma_x}}}$$
$$= \frac{\bar{x} - \mu_x}{(s_x/\sqrt{n})}$$

- Note the similarity to the z-score: the only difference here is that we estimate σ_x with s_x
- Later we will use the t-distribution when we make inference on the mean and don't know the population standard deviation

$$t_{n-1} = \frac{\bar{x} - \mu_x}{(s_x/\sqrt{n})}$$

 This result is something we will use in Chapter 11 and deals with the relationship between the F and the t distribution.

• The F density can be defined as follows:

$$f_{x(x)} = \frac{\left(\sqrt{\frac{(d_1 x)^2 d_2^2}{(d_1 x + d_2)^{(d_1 + d_2)}}}\right)}{xB\left(\frac{d_1}{2}, \frac{d_2}{2}\right)} I(x \ge 0)$$

- Mean = $\frac{d_2}{d_2 2}$ for $d_2 > 2$
- Variance = $\frac{2*d_2^2(d_1+d_2-2)}{d_1(d_2-2)^2(d_2-4)}$ for $d_2 > 4$

P(X = x) = 0 as the probability of any one value is always zero

•
$$P(X \leq x) = pf(x, n_x, n_y)$$

•
$$P(X \ge x) = 1 - pf(x, n_x, n_y)$$

• $P(x_1 < X < x_2) = pf(x_2, n_x, n_y) - pf(x_1, n_x, n_y)$



$$n_x = 2, n_y = 1$$



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• The F distribution can be defined as follows:

•
$$F_{n_x,n_y} = \frac{\left(\frac{X_x^2}{n_x}\right)}{\left(\frac{X_y^2}{n_y}\right)}$$

• Where X_x^2 with n_x degrees of freedom and X_y^2 with n_y degrees of freedom are independent χ^2 random variables

• Relating it back to \bar{x} , s_x^2 , and the t-distribution

$$F_{n_x-1,n_y-1} = \frac{\left(\frac{\left(\frac{(n_x-1)s_x^2}{\sigma_x^2}\right)}{n_x-1}\right)}{\left(\frac{\left(\frac{(n_y-1)s_y^2}{\sigma_y^2}\right)}{n_y-1}\right)} = \frac{\frac{s_x^2}{\sigma_x^2}}{\frac{s_y^2}{\sigma_y^2}} = \frac{\frac{s_x^2}{\sigma_y^2}}{\frac{\sigma_x^2}{\sigma_y^2}}$$

 Thus allowing us to compare the variances of two populations

• We say
$$\frac{\left(\frac{s_x^2}{s_y^2}\right)}{\left(\frac{\sigma_x^2}{\sigma_y^2}\right)}$$
 follows F_{n_x-1,n_y-1} , the F distribution with $n_x - 1$ and $n_y - 1$ degrees of freedom

Summaries!

• Bunnies, Rabbits and the NY times

– <u>https://www.youtube.com/watch?v=jvoxEYmQHNM</u>

Sampling Distribution for the Sample Proportion Summary

Shape of sample	Center of sample	Spread of sample
The shape of the distribution is bell shaped if $n * \rho \ge 15$ and $n * (1 - \rho) \ge 15$	$\mu_{\hat{p}} = \rho$	$\sigma_{\hat{p}} = \sqrt{\frac{\rho(1-\rho)}{n}}$

Sampling Distribution for the Sample Mean Summary

Shape, Center and Spread of Population	Shape of sample	Center of sample	Spread of sample
Population is normal with mean μ and standard deviation σ .	Regardless of the sample size n, the shape of the distribution of the sample mean is normal	$\mu_{\bar{x}} = \mu$	$\sigma_{\bar{\chi}} = \frac{\sigma_{\chi}}{\sqrt{n}}$
Population is not normal with mean μ and standard deviation σ .	As the sample size n increases, the distribution of the sample mean becomes approximately normal	$\mu_{\bar{x}} = \mu$	$\sigma_{\bar{x}} = \frac{\sigma_x}{\sqrt{n}}$

s_{χ}^2 and the χ^2 (chi-squared) distribution

• The chi-squared density can be defined as follows:

$$f_{x(x)} = \frac{1}{2^{\frac{n}{2}}\Gamma\left(\frac{n}{2}\right)} x^{\frac{n}{2}-1} e^{-\frac{x}{2}} I(x > 0)$$

- Mean = n
- Variance = 2n

χ^2 (chi-squared) distribution in R

• P(X = x) = 0 as the probability of any one value is always zero

- $P(X \le x) = \text{pchisq}(x, n)$
- $P(X \ge x) = 1 \text{pchisq}(x, n)$
- $P(x_1 < X < x_2) = pchisq(x_2, n) pchisq(x_1, n)$

s_{χ}^2 and the χ^2 (chi-squared) distribution

• The chi-squared distribution can be defined as follows:

$$X^2 = z_1^2 + z_2^2 + \dots + z_n^2$$

 X² follows the chi-squared distribution with n degrees of freedom where z_i are independent random variables that follow the standard normal distribution

The Chi-Squared Distribution

$$X_{n-1}^2 = \left(\frac{(n-1)s^2}{\sigma_x^2}\right) = \sum \left(\frac{x_i - \bar{x}}{\sigma_x}\right)^2$$

• $\left(\frac{(n-1)s^2}{\sigma_x^2}\right)$ follows a χ^2 distribution with n-1 degrees of freedom

• The t density can be defined as follows:

$$f_{x(x)} = \frac{\Gamma\left(\frac{n+1}{2}\right)}{\sqrt{\pi n}\Gamma\left(\frac{n}{2}\right)} \left(1 + \frac{x^2}{n}\right)^{-\frac{n+1}{2}} I(X \in \mathbb{R})$$

- Mean = 0
- Variance = $\frac{n}{n-2}$

• P(X = x) = 0 as the probability of any one value is always zero

•
$$P(X \le x) = pt(x, n)$$

- $P(X \ge x) = 1 pt(x, n)$
- $P(x_1 < X < x_2) = pt(x_2, n) pt(x_1, n)$

• The t distribution can be defined as follows:

$$t_n = \frac{Z}{\sqrt{\frac{\chi_n^2}{n}}}$$

• t_n follows the t distribution with n degrees of freedom where Z follows the standard normal distribution and χ^2 is chi-squared and divided by its degrees of freedom

$$t_{n-1} = \frac{Z}{\sqrt{\frac{\chi_n^2}{n}}} = \frac{\bar{x} - \mu_x}{(s_x/\sqrt{n})}$$

- $\frac{\bar{x} \mu_x}{(s_x/\sqrt{n})}$ follows a t distribution with n-1 degrees of freedom
- Note the similarity to the z-score: the only difference here is that we estimate σ_x with s_x

• The F density can be defined as follows:

$$f_{x(x)} = \frac{\left(\sqrt{\frac{(d_1 x)^2 d_2^2}{(d_1 x + d_2)^{(d_1 + d_2)}}}\right)}{xB\left(\frac{d_1}{2}, \frac{d_2}{2}\right)} I(x \ge 0)$$

- Mean = $\frac{d_2}{d_2 2}$ for $d_2 > 2$
- Variance = $\frac{2*d_2^2(d_1+d_2-2)}{d_1(d_2-2)^2(d_2-4)}$ for $d_2 > 4$

P(X = x) = 0 as the probability of any one value is always zero

•
$$P(X \leq x) = pf(x, n_x, n_y)$$

•
$$P(X \ge x) = 1 - pf(x, n_x, n_y)$$

• $P(x_1 < X < x_2) = pf(x_2, n_x, n_y) - pf(x_1, n_x, n_y)$

• The F distribution can be defined as follows:

•
$$F_{n_x,n_y} = \frac{\left(\frac{X_x^2}{n_x}\right)}{\left(\frac{X_y^2}{n_y}\right)}$$

• Where X_x^2 with n_x degrees of freedom and X_y^2 with n_y degrees of freedom are independent χ^2 random variables

$$F_{n_x-1,n_y-1} = \frac{\left(\frac{S_x^2}{S_y^2}\right)}{\left(\frac{\sigma_x^2}{\sigma_y^2}\right)}$$

$$\frac{\left(\frac{s_x^2}{s_y^2}\right)}{\left(\frac{\sigma_x^2}{\sigma_y^2}\right)} \text{ follows, the F distribution with } n_x - 1 \text{ and } n_y - 1 \text{ degrees of freedom}$$

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